

MATHEMATICAL TRIPOS Part III

Wednesday 1 June, 2016 $\,$ 9:00 am to 12:00 pm $\,$

PAPER 124

PROBABILISTIC NUMBER THEORY

Attempt no more than **THREE** questions. There are **FIVE** questions in total. The questions carry equal weight.

STATIONERY REQUIREMENTS

Cover sheet Treasury Tag Script paper **SPECIAL REQUIREMENTS** None

You may not start to read the questions printed on the subsequent pages until instructed to do so by the Invigilator.

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(a) Define an additive function and a completely additive function.

Prove that if f is an additive function, and if \mathbb{E}_N denotes expectation with respect to the discrete uniform probability measure \mathbb{P}_N on $[N] := \{1 \leq n \leq N\}$, then

$$\mathbb{E}_N f = \sum_{p^k \leqslant N} \frac{f(p^k)}{p^k} \left(1 - \frac{1}{p} \right) + O\left(\frac{1}{N} \sum_{p^k \leqslant N} |f(p^k)| \right).$$

Let $\omega(n)$ and $\Omega(n)$ denote the number of prime factors of n counted without multiplicity and with multiplicity, respectively. Prove that if t(N) is any function that tends to infinity as $N \to \infty$, then

$$\mathbb{P}_N(\Omega - \omega \ge t(N)) \to 0 \text{ as } N \to \infty.$$

(b) Define the Möbius function $\mu(n)$, and define the Riemann zeta function $\zeta(s)$ for $\Re(s) > 1$.

Prove carefully that when $\Re(s) > 1$ we have

$$\zeta(s) = \prod_{\text{primes } p} \left(1 - \frac{1}{p^s}\right)^{-1}, \text{ and also } \zeta(s) \cdot \sum_{n=1}^{\infty} \frac{\mu(n)}{n^s} = 1.$$

(c) Suppose we knew that for all small $\epsilon > 0$, we had $\sum_{n \leq x} \mu(n) \ll_{\epsilon} x^{0.9+\epsilon}$ for all x. Deduce the strongest statement you can about the zeros of $\zeta(s)$.

[You may assume that $\zeta(s)$ can be analytically continued to a meromorphic function on \mathbb{C} , with only a simple pole at s = 1, and you may assume any other basic facts about the zeta function provided you state them clearly.]

CAMBRIDGE

 $\mathbf{2}$

(a) Prove the First Borel–Cantelli Lemma, which states that if $(A_n)_{n=1}^{\infty}$ is a sequence of events (measurable with respect to a probability measure \mathbb{P}), and if $\sum_{n=1}^{\infty} \mathbb{P}(A_n)$ converges, then

 $\mathbb{P}(\text{infinitely many of the } A_n \text{ occur}) =: \mathbb{P}(A_n \text{ i.o.}) = 0.$

(b) Let $(U_n)_{n=1}^{\infty}$ be as in the Cramér model, so that $U_1 = 0$, and $U_2 = 1$, and $(U_n)_{n \ge 3}$ is a sequence of independent Bernoulli random variables taking value 1 with probability $1/\log n$, and taking value 0 otherwise.

For each $n \ge 1$, define random variables $P_n := \min\{m : \sum_{i=1}^m U_i = n\}$. Prove that

$$\mathbb{P}(\limsup_{n \to \infty} \frac{P_{n+1} - P_n}{\log^2(P_n)} \leqslant 1) = 1.$$

(c) Let $S_n = \sum_{i=1}^n X_i$ be a simple random walk, so that X_i are independent random variables taking values ± 1 with probability 1/2 each. Prove that for any $\theta \in \mathbb{R}$ we have

$$\mathbb{E}\exp\{\theta S_n\} = \left(\frac{e^{\theta} + e^{-\theta}}{2}\right)^n \leqslant e^{n\theta^2/2}.$$

Using this result, together with the fact (which you may assume without proof) that for any $\theta > 0$ and $t \in \mathbb{R}$ we have

$$\mathbb{P}(\max_{k \leq n} S_k \geqslant t\sqrt{n}) \leqslant e^{-\theta t\sqrt{n}} \mathbb{E} \exp\{\theta S_n\},\$$

deduce that for any $t \ge 0$ we have

$$\mathbb{P}(\max_{k \leq n} S_k \geqslant t\sqrt{n}) \leq e^{-t^2/2}.$$

By combining this with part (a), prove that

$$\mathbb{P}(\limsup_{n \to \infty} \frac{S_n}{\sqrt{2n \log \log n}} \leqslant 1) = 1.$$

[Hint: You may find it helpful to consider the special sequence of values $n = \lfloor (1+\delta)^m \rfloor$, for small fixed $\delta > 0$ and $m \in \mathbb{N}$.]

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(a) Define a *z*-sieved number n, where $z \ge 2$.

Define the Buchstab function w(u) for $u \ge 1$, and state Buchstab's theorem about $\Phi(z^u, z) := \#\{n \le z^u : n \text{ is } z - \text{sieved}\}.$

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(b) You may assume that $\lim_{u\to\infty} w(u) = e^{-\gamma}$, where $\gamma \approx 0.577$ is Euler's constant. You may also assume that for all $u \ge 1$ we have $|w(u) - e^{-\gamma}| \ll \frac{1}{\Gamma(u+1)}$, where $\Gamma(u+1) = \int_0^\infty e^{-x} x^u dx$ is the classical gamma function.

Prove that for any $u \ge 1$, there exists a point $u \le u^+ \le u+1$ at which $w(u^+) > e^{-\gamma}$ and there exists a point $u \le u^- \le u+1$ at which $w(u^-) < e^{-\gamma}$.

(c) Let N be large, let \mathbb{P}_N denote the discrete uniform probability measure on [N], and let $\omega(n)$ denote the number of distinct prime factors of n.

State the Turán–Kubilius inequality for the variance of an additive function, and use it to show that

$$\mathbb{P}_N(|\omega(n) - \log \log N| \ge 0.01 \log \log N) \ll \frac{1}{\log \log N}.$$

[You may use any standard estimates for sums over primes provided you state them clearly, but you should briefly prove any other estimates that you need.]

Using the above estimate (and NOT using any estimate involving the function $\Omega(n)$), obtain the best upper bound you can for

$$#\{ab: 1 \leq a \leq N, \ 1 \leq b \leq N\}.$$

 $\mathbf{4}$

(a) Let $A(s) = \sum_{n \leq X} \frac{a_n}{n^s}$ and $B(s) = \sum_{n \leq X} \frac{b_n}{n^s}$, where a_n, b_n are arbitrary complex numbers. Prove that for any $T \ge 0$ and $\sigma \in \mathbb{R}$, we have

$$\int_{T}^{2T} A(\sigma + it) \overline{B(\sigma + it)} dt = T \sum_{n \leqslant X} \frac{a_n \overline{b_n}}{n^{2\sigma}} + O\left(\sum_{m,n \leqslant X, m \neq n} \frac{|a_n| |b_m|}{n^{\sigma} m^{\sigma} |\log(m/n)|}\right)$$

Show that the "big Oh" term here is $\ll X \sum_{m,n \leq X, m \neq n} \frac{|a_n||b_m|}{n^{\sigma}m^{\sigma}}$, and show also that it is

$$\ll \sum_{n \leqslant X} \frac{|a_n|^2 n \log(2X)}{n^{2\sigma}} + \sum_{m \leqslant X} \frac{|b_m|^2 m \log(2X)}{m^{2\sigma}}.$$

(b) Prove that for any $X \ge 2$, any $T \ge 0$, any $\sigma \in \mathbb{R}$, and any odd $j \in \mathbb{N}$, we have

$$\int_{T}^{2T} \left(\sum_{\text{primes } p \leqslant X} \frac{\cos(t \log p)}{p^{\sigma}} \right)^{j} dt \ll_{j} X^{j} (\sum_{n \leqslant X^{j}} \frac{1}{n^{\sigma}})^{2}.$$

(c) You may assume that for all $t \ge 1$ we have

$$|\zeta(1/2+it)| \ll |\sum_{n \leqslant \sqrt{t/2\pi}} \frac{1}{n^{1/2+it}}| + \frac{1}{t^{1/4}},$$

where $\zeta(s)$ denotes the Riemann zeta function (a meromorphic function with only a simple pole at s = 1).

You may also assume that for all $X \ge 2$ we have $\sum_{n \le X} d(n)^2 \ll X(\log X)^3$, where d(n) denotes the total number of divisors of n.

By using these facts, and suitably adapting the argument in part (a), prove that for all $T \ge 2$ we have

$$\int_0^T |\zeta(1/2 + it)|^4 dt \ll T \sum_{\substack{n_1, \dots, n_4 \le \sqrt{T/2\pi}, \\ n_1n_2 = n_3n_4}} \frac{1}{\sqrt{n_1 n_2 n_3 n_4}} + T(\log T)^4.$$

Deduce that overall we have

$$\int_0^T |\zeta(1/2 + it)|^4 dt \ll T(\log T)^4.$$

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[TURN OVER

 $\mathbf{5}$

(a) State a general form of the Erdős–Kac central limit theorem.

(b) Suppose g(n) is a strongly additive function (that is, an additive function for which $g(p^k) = g(p)$ for all $k \in \mathbb{N}$). Suppose also that N is large, that $\phi(N) \ge 1$ is some function, and that g(p) = 0 whenever $p > N^{1/\phi(N)}$.

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Furthermore, let $(X_q)_{q \text{ prime}}$ be a sequence of independent random variables (measurable with respect to some probability measure \mathbb{P} with corresponding expectation \mathbb{E}), such that X_q takes value 1 with probability 1/q, and it takes value 0 otherwise.

Prove that if \mathbb{E}_N denotes expectation with respect to the discrete uniform probability measure \mathbb{P}_N on [N], then for any natural number j we have

$$\mathbb{E}_{N}(g - \mathbb{E}_{N}g)^{j} = \sum_{p_{1},\dots,p_{j} \leqslant N^{1/\phi(N)}} \prod_{i=1}^{w} g(q_{i})^{a_{i}} \mathbb{E}(X_{q_{i}} - \frac{1}{q_{i}})^{a_{i}} + O\Big(\frac{jN^{j/\phi(N)}}{N}\Big(\sum_{p \leqslant N^{1/\phi(N)}} |g(p)|\Big)^{j}\Big).$$

Here in the sum we write q_i for the *distinct* primes amongst $p_1, ..., p_j$, and a_i for the multiplicities with which they occur, so always $a_1 + ... + a_w = j$.

(c) State a version of the *method of moments* for proving convergence in distribution.

Explain *briefly* how this may be used to prove the Erdős–Kac theorem.

[You should remark briefly on the role of part (b) in the proof, and on the choice of the function $\phi(N)$, but you do NOT need to give any technical details.]

(d) Let $\omega(n)$ denote the number of distinct prime factors of n, and for all large N define

$$E(N) := \sup_{z \in \mathbb{R}} \left| \mathbb{P}_N(\frac{\omega - \log \log N}{\sqrt{\log \log N}} \leqslant z) - \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-t^2/2} dt \right|.$$

Prove the lower bound

$$E(N) \gg \frac{1}{\sqrt{\log \log N}}.$$

END OF PAPER